

The Global Rise of Asset Prices and the Decline of the Labor Share

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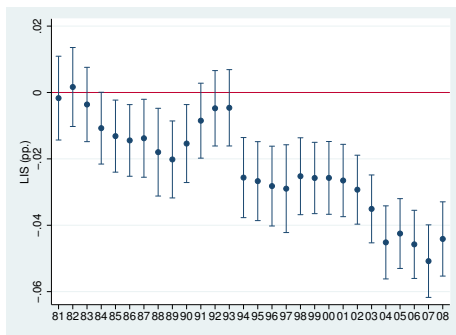
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Motivation: Fact#1

- The global decline of the labor income share since the 1980s has sparked new interest in the functional distribution of income.

Figure 1: Labor income share, 1980-2008



Notes: Own calculations obtained as year fixed effects from a GDP weighted regression including country fixed effects to control for the entry and exit of countries throughout the sample.

Motivation

- Attempts to explain this trend: The role of the capital-output ratio (k)
 - ▶ Bentolila and Saint-Paul (2003) CES technology $\Rightarrow lis = f(\frac{k}{y})$
 - ▶ Piketty and Zucman (2014) $s > g \Rightarrow \uparrow \frac{k}{y} \Rightarrow \downarrow lis$ if $\sigma > 1$
 - ▶ Karabarbounis and Neiman (2014) $\downarrow rp \Rightarrow \uparrow \frac{k}{y} \Rightarrow \downarrow lis$ if $\sigma > 1$
 - ▶ Koh *et al* (2016) $\uparrow k_{IPP} \Rightarrow \uparrow \frac{k}{y} \Rightarrow \downarrow lis$ if $\sigma > 1$
- Secular increases in the capital-output ratio are the main cause of the long-run labor share decline. [Graph](#)
- This requires $\sigma > 1$, a value which has seldom been found in the literature ([Chirinko and Mallick, 2014](#)).

Our Contribution

- Finance and assets prices have developed since 1980s ([Philippon, 2012](#); [Greenwod and Scharfstein, 2013](#)).
- The “accumulation view” has ignored the role of asset prices (although [Piketty and Zucman \(2014\)](#) document that part of the rise of wealth ratios comes from changes in asset prices).
- In this paper we explore a new mechanism that connects the rise of stock prices with the decline of the labor share (through lower capital-deepening).
- Panel Time Series.
- We find evidence for several mechanisms that operate through the same channel: i) the rise of monopoly mark-ups, ii) the decline of dividend income taxes and iii) the rise of corporate short-termism.

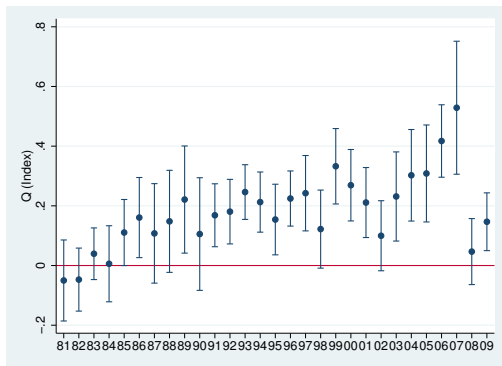
Outline

- ① Facts
- ② Theoretical Framework
- ③ Data
- ④ Empirical Methodology
- ⑤ Results
- ⑥ Conclusions

Fact#2

- Steady increase in financial wealth with respect to productive capital.

Figure 2: Tobin's Q , 1980-2009



Notes: Own calculations obtained as year fixed effects from a GDP weighted regression including country fixed effects to control for the entry and exit of countries throughout the sample.

Theoretical Framework: Households

- Representative household accumulates financial assets.
- Direct utility from the ownership of wealth ($ps' = a'$) like in Carroll (1998), Piketty (2011) and Saez and Stantcheva (2017).

$$\begin{aligned} U(a) &= \max_{c, a'} u(c) + h(a) + \beta U(a') \\ \text{s.t. } & c + a' = w + (1 + r)a, \end{aligned} \tag{1}$$

- where $a' = ps'$
- and $1 + r = \frac{(1-\tau)d+p}{p-1}$
- The demand of assets $a(r)$ is an increasing function.

Theoretical Framework: Firms

- CES Technology

$$y = \left[\phi k^{\left(\frac{\sigma-1}{\sigma}\right)} + (1 - \phi) l^{\left(\frac{\sigma-1}{\sigma}\right)} \right]^{\frac{\sigma}{\sigma-1}}, \quad (2)$$

- Monopolistic firms (elasticity for each variety is ξ) that maximize their market value, accumulate physical capital and distribute dividends to households.
- Symmetric FOC wrt to k'

$$F_k(k, l) = \left(\frac{\xi}{\xi - 1} \right) (\delta + r), \quad (3)$$

Market

- Tobin's Q (at the steady state):

$$Q(r) = (1 - \tau) \left(1 + \frac{F(k(r), l)}{\xi r k(r)} \right)$$

- Equity Wealth (at the steady state):

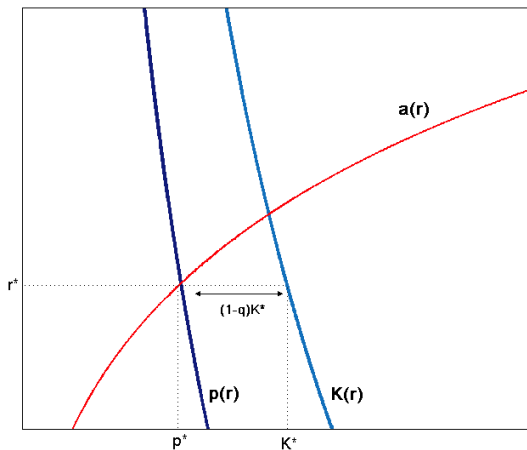
$$p(r) = Q(r)k(r) = (1 - \tau) \left(k(r) + \frac{F(k(r), l)}{\xi r} \right)$$

- Market clearing

$$a(r^*) = p(r^*) \equiv Q(r^* | \tau, \xi) k(r^*)$$

Asset Prices and Productive Capital

Figure 3: Market for capital



Impact on the Labor Share

- Assuming a CES production function, we have:

$$LIS = 1 - \phi \left(\frac{k}{y} \right)^{\frac{\sigma-1}{\sigma}}$$

- From our model, equilibrium relation between Tobin's Q and LIS :

$$LIS^*(Q) = 1 - \phi \left(\frac{k^*(Q)}{y^*(Q)} \right)^{\frac{\sigma-1}{\sigma}} \quad \text{where} \quad \frac{\partial LIS}{\partial \frac{K}{Y}} > 0 \quad \text{if } \sigma < 1; \quad (4)$$

- Therefore:

$$\frac{\partial LIS}{\partial Q} = \left(\frac{\partial LIS}{\partial \frac{k}{y}} \right) \left(\frac{\partial \frac{k}{y}}{\partial k} \right) \left(\frac{\partial k(Q)}{\partial Q} \right) < 0 \quad (5)$$

because:

$$\frac{\partial \frac{k}{y}}{\partial k} > 0 \quad \text{and} \quad \frac{\partial k(Q)}{\partial Q} < 0$$

Data

41 countries, 1980-2009. [Sample](#)

- Tobin's Q :
 - ▶ Worldscope Database.
 - ▶ [Doidge *et al.* \(2013\)](#) methodology.
- Labor income share:
 - ▶ Extended Penn World Table 4.0.
 - ▶ No adjustment for mixed rents, no distinction of the corporate sector.
 - ▶ Correlation between 0.87 and 0.96 with [Karabarounis and Neiman \(2014\)](#).
- Relative prices:
 - ▶ Extension of [Karabarounis and Neiman \(2014\)](#) database.
 - ▶ Penn World Table 7.1 and BEA.

Empirical Implementation

We assume a general multiplicative form between our variables of interest:

$$LIS = g\left(\frac{k}{y}\right) = a\left(\frac{k}{y}\right)^\alpha, \quad \text{and} \quad \frac{k}{y} = f(Q, RP) = Q^{\psi_1} RP^{\psi_2} \quad (6)$$

We use these two forms to obtain an estimable equation of the labor share in terms of Q and RP :

$$LIS = g\left(\frac{k}{y}\right) = g(f(Q, RP)) = a(Q^{\psi_1} RP^{\psi_2})^\alpha \quad (7)$$

or in logs:

$$lis_{it} = \beta_0 + \beta_1 q_{it} + \beta_2 rp_{it} + \Omega_{it} \quad (8)$$

Empirical Methodology

- Macroeconomics panel data makes difficult the use of traditional panel data techniques:
 - ▶ Small N compared to T.
 - ▶ Parameter heterogeneity.
 - ▶ Cross-section dependence. Pesaran (2004) CD test
 - ▶ Nonstationary data. Pesaran (2007) CIPS test
- New Panel Time Series Techniques based on Common factor models:

$$y_{it} = \beta_i x_{it} + u_{it}, \quad u_{it} = \varphi_i f_t + \psi_i + \varepsilon_{it}, \quad (9)$$

$$x_{it} = \delta_i f_t + \pi_i + e_{it}, \quad f_t = \tau + \phi f_{t-1} + \omega_t, \quad (10)$$

where (f_t) represents unobserved time-variant heterogeneity and raises endogeneity problems which make difficult the estimation of β_i .

Empirical Methodology

- Common Correlated Effect estimators: Observed regressors are augmented with cross-sectional averages of the dependent variable and the individual-specific regressors ([Pesaran, 2006](#)). Intuition

$$\begin{aligned}lis_{it} &= \beta_{i0} + \beta_{i1}q_{it} + \beta_{i2}rp_{it} \\ &+ \beta_{i3}\overline{lis_{it}} + \beta_{i4}\overline{q_{it}} + \beta_{i5}\overline{rp_{it}} + \Omega_{it}\end{aligned}$$

- Our reference results are obtained from an ECM version using the [Chudik and Pesaran \(2015\)](#) Dynamic Common Correlated Effects Mean Group Estimator:

$$\begin{aligned}\Delta lis_{it} &= \beta_{i0} + \beta_{i1}lis_{i,t-1} + \beta_{i2}q_{i,t-1} + \beta_{i3}rp_{i,t-1} + \beta_{i4}\Delta q_{it} + \beta_{i5}\Delta rp_{it} \\ &+ \beta_{i6}\overline{\Delta lis_{it}} + \beta_{i7}\overline{lis_{i,t-1}} + \beta_{i8}\overline{q_{i,t-1}} + \beta_{i9}\overline{rp_{i,t-1}} + \beta_{i10}\overline{\Delta q_{it}} + \beta_{i11}\overline{\Delta rp_{it}} \\ &+ \sum_{l=1}^p \beta_{i12}\overline{\Delta lis_{i,t-l}} + \sum_{l=1}^p \beta_{i13}\overline{\Delta q_{i,t-l}} + \sum_{l=1}^p \beta_{i14}\overline{\Delta rp_{i,t-l}} + \Omega_{it},\end{aligned}$$

Table 1: Results: Error Correction Model

| | [1] 2FE | [2] CCEP | [3] MG | [4] CMG | [5] CMGt | [6] CMGt1 | [7] CMGt2 |
|--------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| lis_{t-1} | -0.176 (0.026)*** | -0.395 (0.049)*** | -0.449 (0.034)*** | -0.5 (0.053)*** | -0.694 (0.061)*** | -0.72 (0.085)*** | -0.812 (0.125)*** |
| qt_{t-1} | 0.011 (0.013) | -0.012 (0.015) | -0.035 (0.014)** | -0.039 (0.018)** | -0.067 (0.026)** | -0.076 (0.028)*** | -0.058 (0.033)* |
| rpt_{t-1} | -0.032 (0.024) | -0.016 (0.040) | 0.064 (0.070) | 0.15 (0.091)* | 0.092 (0.115) | 0.129 (0.166) | -0.005 (0.186) |
| Δq | -0.031 (0.014)** | -0.033 (0.015)** | -0.038 (0.009)*** | -0.038 (0.012)*** | -0.051 (0.017)*** | -0.053 (0.019)*** | -0.058 (0.018)*** |
| Δrp | -0.141 (0.050)*** | -0.214 (0.056)*** | -0.021 (0.065) | 0.049 (0.108) | 0.093 (0.099) | 0.05 (0.107) | -0.11 (0.095) |
| t | | | 0.001 (0.001) | | 0.001 (0.002) | 0.001 (0.003) | 0.001 (0.004) |
| Constant | -0.106 (0.018)*** | | -0.301 (0.033)*** | -0.273 (0.050)*** | -0.277 (0.084)*** | -0.431 (0.089)*** | -0.356 (0.124)*** |
| Number of id | 30 | 30 | 30 | 30 | 30 | 29 | 26 |
| Observations | 732 | 732 | 732 | 732 | 732 | 700 | 631 |
| R-squared | 0.26 | 0.59 | | | | | |
| RMSE | 0.0264 | 0.0224 | 0.0191 | 0.0142 | 0.0127 | 0.0101 | 0.0067 |
| Trends | | | 0.23 | | 0.20 | 0.21 | 0.23 |
| lr- q | 0.0621 | -0.0307 | -0.0779 | -0.0785 | -0.0965 | -0.1061 | -0.0718 |
| se- q | 0.0739 | 0.0357 | 0.0327 | 0.0374 | 0.0388 | 0.0405 | 0.0422 |
| lr- rp | -0.1826 | -0.0405 | 0.1417 | 0.2999 | 0.1325 | 0.1796 | -0.0063 |
| se- rp | 0.1306 | 0.1016 | 0.1573 | 0.185 | 0.1661 | 0.2312 | 0.2285 |
| CD test | -2.4749 | -1.5637 | 4.9547 | -0.0134 | -0.2654 | 1.0079 | 1.3218 |
| Abs Corr | 0.1884 | 0.217 | 0.2038 | 0.2189 | 0.2216 | 0.2393 | 0.2466 |
| Int | I(0) | I(0) | I(0) | I(0) | I(0) | I(0) | I(0) |

Robustness

Weak Exogeneity Test

Table 1: Results: Error Correction Model

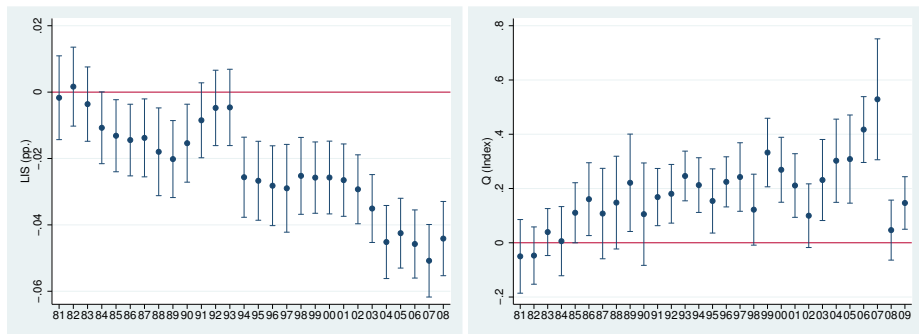
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| t | | | 0.001 (0.001) | | 0.001 (0.002) | 0.001 (0.003) | 0.001 (0.004) |
| Constant | -0.106 (0.018)*** | | -0.301 (0.033)*** | -0.273 (0.050)*** | -0.277 (0.084)*** | -0.431 (0.089)*** | -0.356 (0.124)*** |
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| $se-q$ | 0.0739 | 0.0357 | 0.0327 | 0.0374 | 0.0388 | 0.0405 | 0.0422 |
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| Int | I(0) | I(0) | I(0) | I(0) | I(0) | I(0) | I(0) |

Robustness

Weak Exogeneity Test

Quantification

Figure 4: Labor income share and Tobin's Q , 1980-2009



(a) Labor Income Share

(b) Tobin's Q

- Tobin's Q has increased (52%) from 1.15 (1980) to 1.68 (2007).
- LIS has evolved from 57% (1980) to 52% (2007) (-8.9%).
- Tobin's Q can explain between 41% and 57%.

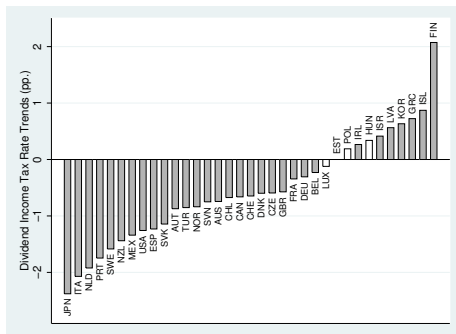
Beyond the Q : What is Behind?

- Dividend Income Tax Rate
- Market Power: The Industry Concentration Rate
- Corporate Governance

Dividend Income Tax Rate (I)

- Data (Max 1980-2014):
 - ▶ Dividend Income Tax Rate: OECD Tax Database
 - ▶ Capital-Output Ratio: AMECO
 - ▶ Tobin's Q : Worldscope

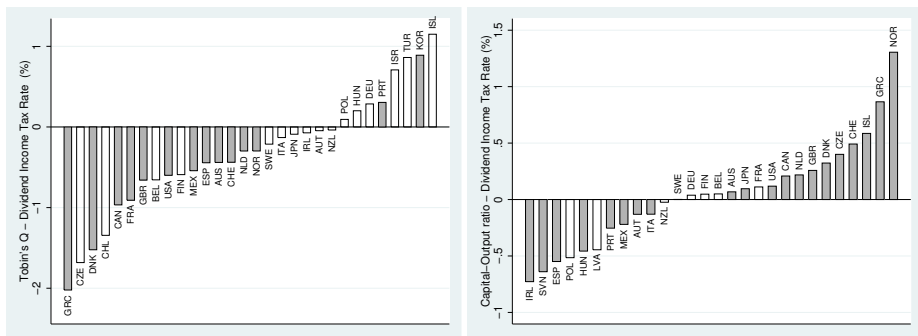
Figure 5: Country-specific Trends: Dividend Tax Rate



Notes: Own calculations obtained from $TAX_t = \alpha_0 + \alpha_1 t + \epsilon_t$, where Tax is the dividend tax rate, t is a linear trend, and \epsilonpsilon is a classic disturbance term. The vertical axis show α_1 in %. Dark bars indicate that α_1 is significant at 5% level. Each regression only includes countries which have at least 10 observations for the period 1980-2014.

Dividend Income Tax Rate (II)

Figure 6: Tobins' Q, Capital-Output Ratios and Dividend Income Tax Rates



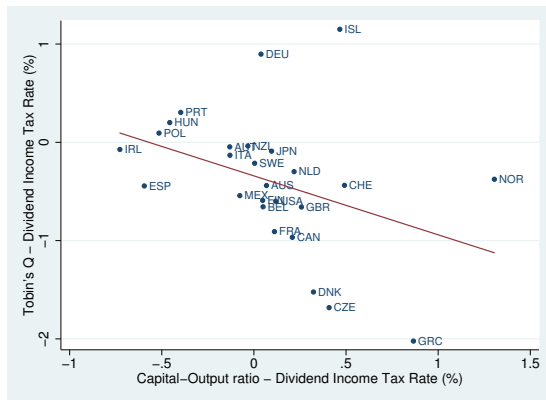
(a) Tobins' Q

(b) Capital-Output ratio

Notes: Own calculations obtained from $\ln(X_t) = \alpha_0 + \alpha_1 TAX_t + \epsilon_t$, where X represents the Tobin's Q or the capital-output ratio, TAX stands for the dividend income tax rate, and ϵ is a classic disturbance term. The vertical axis shows the coefficient α_1 in %. Dark bars indicate that α_1 is significant at 5% level. Each graph shows countries for which we have at least 10 observations for the period under analysis (Max. 1980-2014). Luxembourg is excluded from the graph due to be a clear outlier.

Dividend Income Tax Rate (III)

Figure 7: Tobin's Q , Capital-Output Ratios and Dividend Income Tax Rates



Notes: Own calculations obtained from $\ln(X_t) = \alpha_0 + \alpha_1 TAX_t + \epsilon_t$, where X represents the Tobin's Q and the capital-output ratio in the vertical and the horizontal axis respectively. TAX is the dividend income tax rate, and ϵ is a classic disturbance term. Both axis show the coefficient α_1 in %. Both equations are constraint to have the same number of observations (Max. 1980-2014). The scatter plot is obtained after excluding outliers. An outlier is defined as an observation with a weight of 0 after using the *rrreg* command in STATA.

Market Power: The Industry Concentration Rate (I)

- Data (U.S. Industry Data, 2002-2012):
 - ▶ Market Power: U.S. Economic Census
 - ▶ Capital-Output Ratio: NBER-CES Manufacturing Industry
 - ▶ Tobin's Q: Worldscope

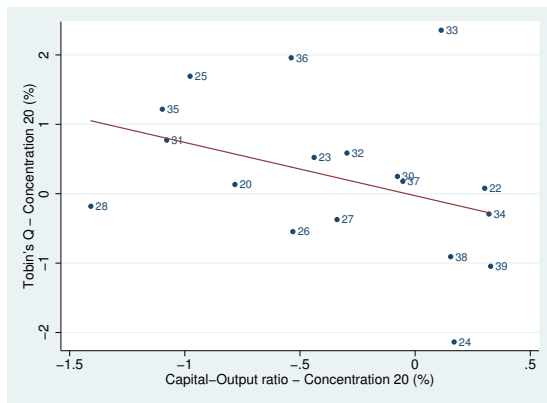
Table 2: Tobin's Q, Capital-Output Ratio and Industry Concentration

| | [1] | [2] | [3] | [4] | [5] | [6] | [7] | [8] | [9] | [10] | [11] | [12] |
|----------------|--------------------------------|--------------------|---------------------|--------------------|---------------------|---------------------|---------------------------------|---------------------|----------------------|----------------------|----------------------|----------------------|
| | Dependent variable: Δq | | | | | | Dependent variable: Δky | | | | | |
| $\Delta Con4$ | 0.066 (0.078) | | | | 0.087 (0.083) | | | | -0.153 (0.068)** | | | |
| $\Delta Con8$ | | 0.088 (0.109) | | | | 0.120 (0.118) | | | | -0.172 (0.087)* | | |
| $\Delta Con20$ | | | 0.271 (0.126)** | | | | 0.332 (0.134)** | | | | | -0.160 (0.097) |
| $\Delta Con50$ | | | | 0.340 (0.157)** | | | | 0.413 (0.174)** | | | | -0.099 (0.094) |
| Constant | 0.28 (0.031)*** | 0.28 (0.030)*** | 0.278 (0.030)*** | 0.28 (0.031)*** | 0.315 (0.027)*** | 0.315 (0.027)*** | 0.315 (0.027)*** | 0.317 (0.028)*** | -0.079 (0.014)*** | -0.082 (0.014)*** | -0.083 (0.014)*** | -0.083 (0.015)*** |
| R-squared | 0.11 | 0.11 | 0.12 | 0.12 | 0.16 | 0.16 | 0.17 | 0.17 | 0.26 | 0.26 | 0.25 | 0.25 |
| Observations | 834 | 833 | 832 | 825 | 834 | 833 | 832 | 825 | 467 | 467 | 465 | 458 |
| SIC4 | 480 | 480 | 480 | 473 | 480 | 480 | 480 | 473 | 280 | 280 | 280 | 273 |
| SIC2 | 59 | 59 | 59 | 59 | 59 | 59 | 59 | 59 | 20 | 20 | 20 | 20 |
| Sectors | 6 | 6 | 6 | 6 | 6 | 6 | 6 | 6 | 1 | 1 | 1 | 1 |
| Sector FE | YES | YES | YES | YES | NO | NO | NO | NO | NO | NO | NO | NO |
| SIC2 FE | NO | NO | NO | NO | YES | YES | YES | YES | YES | YES | YES | YES |
| TIME FE | YES | YES | YES | YES | YES | YES | YES | YES | YES | YES | YES | YES |

Notes: Robust standard errors clustered at 2-digit SIC level in parenthesis. * significant at 10%; ** significant at 5%; *** significant at 1%. SIC4 and SIC2 indicate the number of groups included in the regressions classified at the 4 and 2-digit SIC level. Sectors indicates the number of groups included using the broader sector definition.

Market Power: The Industry Concentration Rate (II)

Figure 8: Tobins' Q , Capital-Output Ratios and Industry Concentration



Notes: Own calculations obtained from $\Delta \ln(X_{it}) = \alpha_0 + \alpha_1 \Delta \ln(ConY_{it}) + \epsilon_{it}$, where X represents the Tobin's Q and the capital-output ratio in the vertical and the horizontal axis respectively. $Con20$ is the share of sales of the 20 largest companies in the industry, and ϵ is a classic disturbance term. Both axis show the coefficient α_1 in %. Both equations are constraint to have the same number of observations. The scatter plot is obtained after excluding outliers. An outlier is defined as an observation with a weight of 0 after using the `reg` command in STATA.

Corporate Governance (I)

- Data (U.S. Firm Level and Cross-Country, 2002-2014):
 - ▶ Corporate Governance: Asset4 ESG Database
 - ▶ Investment: Worldscope
 - ▶ Capital-Output Ratio: AMECO
 - ▶ Tobin's Q : Worldscope

Figure 9: Tobin's Q , Investment and Corporate Governance (U.S.)

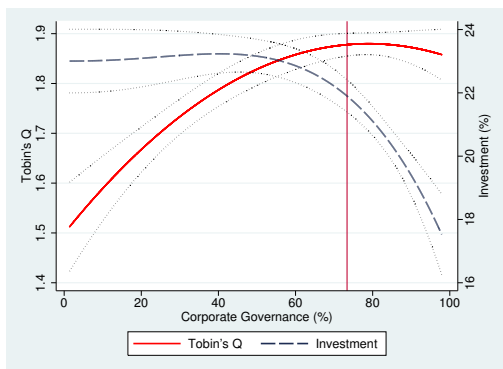


Table 3: Tobin's Q , Investment and Corporate Governance (U.S.)

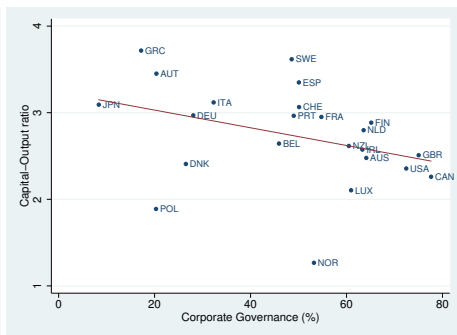
| Panel A | [1] | [2] | [3] | [4] | [5] | [6] | [7] |
|---------------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|----------------------|
| Dependent variable: q | | | | | | | |
| GOV_{t-1} | 0.170 (0.044)*** | 0.178 (0.047)*** | 0.187 (0.048)*** | 0.160 (0.051)*** | 0.083 (0.036)** | 0.151 (0.045)*** | 0.112 (0.040)*** |
| Constant | 0.350 (0.029)*** | 0.446 (0.029)*** | 0.340 (0.032)*** | 0.358 (0.034)*** | 0.409 (0.024)*** | 0.377 (0.031)*** | 0.446 (0.026)*** |
| R-squared | 0.25 | 0.28 | 0.33 | 0.48 | 0.42 | 0.44 | 0.5 |
| Panel B | [1] | [2] | [3] | [4] | [5] | [6] | [7] |
| Dependent variable: INV | | | | | | | |
| GOV_{t-1} | -0.042 (0.019)** | -0.044 (0.019)** | -0.046 (0.019)** | -0.039 (0.018)** | -0.044 (0.018)** | -0.043 (0.017)** | -0.050 (0.019)*** |
| Constant | 0.242 (0.013)*** | 0.218 (0.014)*** | 0.245 (0.013)*** | 0.241 (0.012)*** | 0.244 (0.012)*** | 0.245 (0.011)*** | 0.261 (0.012)*** |
| R-squared | 0.09 | 0.1 | 0.14 | 0.25 | 0.18 | 0.19 | 0.22 |
| Observations | 12574 | 12574 | 12574 | 12574 | 12574 | 12574 | 12574 |
| SIC4 | 365 | 365 | 365 | 365 | 365 | 365 | 365 |
| SIC3 | 212 | 212 | 212 | 212 | 212 | 212 | 212 |
| SIC2 | 62 | 62 | 62 | 62 | 62 | 62 | 62 |
| SIC2 FE | YES | YES | NO | NO | NO | NO | NO |
| SIC3 FE | NO | NO | NO | NO | NO | YES | NO |
| SIC4 FE | NO | NO | NO | NO | YES | NO | YES |
| Time FE | NO | YES | NO | NO | NO | NO | NO |
| SIC2*Time | NO | NO | YES | NO | NO | YES | YES |
| SIC3*Time | NO | NO | NO | YES | NO | NO | NO |

Corporate Governance (III)

Figure 10: Tobin's Q , Capital-Output ratio and Corporate Governance (I)



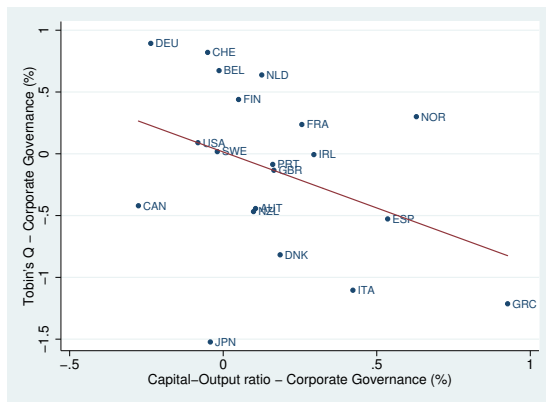
(a) Tobin's Q



(b) Capital-Output ratio

Corporate Governance (IV)

Figure 11: Tobin's Q, Capital-Output ratio and Corporate Governance (II)



Notes: Own calculations obtained from $\ln(X_t) = \alpha_0 + \alpha_1 GOV_t + \epsilon_t$, where X represents the Tobin's Q and the capital-output ratio in the vertical and the horizontal axis respectively. GOV is the corporate governance index, and ϵ is a classic disturbance term. Both axis show the coefficient α_1 in %. Both equations are constraint to have the same number of observations. Each regression only includes countries which have at least 10 observations for the period 2002-2014. The scatter plot is obtained after excluding outliers. An outlier is defined as an observation with a weight of 0 after using the `rreg` command in STATA.

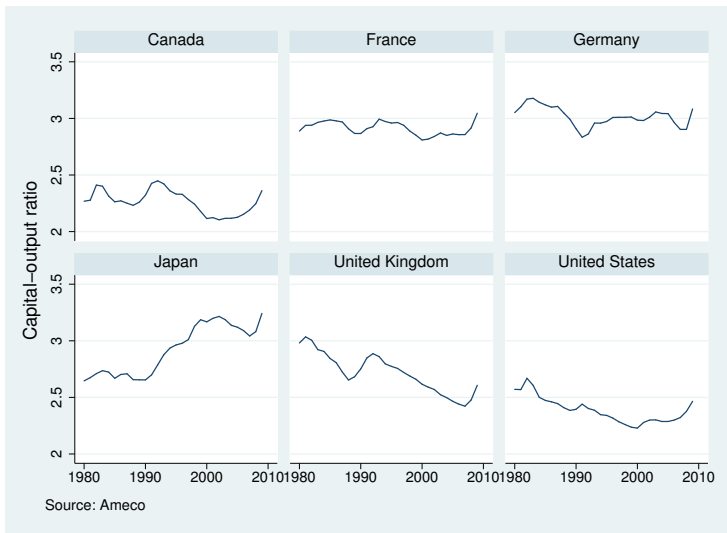
Conclusions

- We find that the increase in Tobin's Q can explain between 41% and 57% of the labor income share decline.
- Our model emphasises the role of asset prices in explaining the decline of the labor share within a standard capital-output framework.
- Relative prices of investment are not relevant.
- Indeed, our model suggests that the problem is not too much physical capital, but the increase of financial wealth with respect to productive capital.
 - ▶ Compatible with standard values of σ (survey [Chirinko, 2008](#)).
- Policies aiming at reversing the trend should target incentives on corporate investment, even if this is at the expense of equity valuation and equity returns.

Additional Materials

Capital-Output ratio

Figure A.1: Capital-Output ratio 1980-2009



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Table A.1: Selected economies and sample period

| id | Country | Sample period | id | Country | Sample period |
|----|-------------|-----------------|----|----------------|---------------|
| 1 | Australia** | 1980-2008 | 22 | Luxembourg* | 1991-2008 |
| 2 | Austria** | 1980-2008 | 23 | Mexico** | 1988-2008 |
| 3 | Belgium** | 1980-2008 | 24 | Morocco | 1998-2007 |
| 4 | Brazil* | 1992-2008 | 25 | Netherlands** | 1980-2008 |
| 5 | Canada** | 1980-2008 | 26 | New Zealand** | 1986-2008 |
| 6 | Chile* | 1990-2008 | 27 | Norway** | 1980-2007 |
| 7 | China | 1995-2007 | 28 | Peru | 1992-2003 |
| 8 | Colombia | 1993-2007 | 29 | Philippines** | 1988-2008 |
| 9 | Denmark** | 1980-2009 | 30 | Poland | 1995-2008 |
| 10 | Finland** | 1987-2009 | 31 | Portugal** | 1988-2009 |
| 11 | France** | 1980-2009 | 32 | South Africa** | 1980-2008 |
| 12 | Germany** | 1983-2008 | 33 | Spain** | 1986-2008 |
| 13 | Greece** | 1988-2009 | 34 | Sri Lanka | 1994-2008 |
| 14 | Hong Kong** | 1980-2003 | 35 | Sweden** | 1982-2009 |
| 15 | Hungary | 1995-2008 | 36 | Switzerland** | 1980-2007 |
| 16 | India* | 1991-2008 | 37 | Thailand | 1988-2003 |
| 17 | Ireland** | 1981-2008 | 38 | Turkey | 1990-2003 |
| 18 | Israel | 1993, 1995-2008 | 39 | UK** | 1980-2008 |
| 19 | Italy** | 1980-2008 | 40 | US** | 1980-2008 |
| 20 | Japan** | 1980-2007 | 41 | Venezuela | 1992-2006 |
| 21 | Korea** | 1980-2003 | | | |

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Table A.2: Cross-section dependence tests

| a) Levels: | | | | b) Diff: | | | |
|-----------------|------------|----------|-----------|-----------------|--------------|------------|-------------|
| Variable | <i>lis</i> | <i>q</i> | <i>rp</i> | Variable | Δlis | Δq | Δrp |
| CD-test | 16.73 | 29.76 | 42.37 | CD-test | 12.99 | 34.45 | 6.66 |
| <i>p</i> -value | 0.00 | 0.00 | 0.00 | <i>p</i> -value | 0.00 | 0.00 | 0.00 |
| corr | 0.132 | 0.250 | 0.345 | corr | 0.11 | 0.296 | 0.049 |
| abs(corr) | 0.472 | 0.394 | 0.558 | abs(corr) | 0.235 | 0.349 | 0.223 |

| c) Het. AR(2) | | | | d) Het. AR(2) CCE | | | |
|-----------------|------------|----------|-----------|-------------------|------------|----------|-----------|
| Variable | <i>lis</i> | <i>q</i> | <i>rp</i> | Variable | <i>lis</i> | <i>q</i> | <i>rp</i> |
| CD-test | 9.93 | 33.58 | 3.40 | CD-test | -0.24 | -0.66 | -2.38 |
| <i>p</i> -value | 0.00 | 0.00 | 0.00 | <i>p</i> -value | 0.81 | 0.51 | 0.02 |
| corr | 0.088 | 0.301 | 0.027 | corr | -0.006 | -0.011 | -0.023 |
| abs(corr) | 0.243 | 0.344 | 0.213 | abs(corr) | 0.220 | 0.237 | 0.213 |

Notes: CD-test shows the Pesaran (2004) cross-section dependence statistic, which follows a $N(0, 1)$ distribution. H_0 = cross-section independence. corr, and abs(corr) report, respectively, the average and average absolute correlation coefficients across the $N(N - 1)$ set of correlations.

Pesaran (2007) CIPS Unit Root Test

Table A.3: Unit root tests

| a) Pesaran (2007) CIPS test: Constant | | | | | | |
|---------------------------------------|------------|--------------|----------|--------------|-----------|--------------|
| Lags | <i>lis</i> | (<i>p</i>) | <i>q</i> | (<i>p</i>) | <i>rp</i> | (<i>p</i>) |
| 0 | 0.431 | 0.667 | -2.744 | 0.003 | -0.118 | 0.453 |
| 1 | -0.207 | 0.418 | -2.405 | 0.008 | -0.141 | 0.444 |
| 2 | -1.199 | 0.115 | 0.103 | 0.541 | 0.655 | 0.744 |
| 3 | 1.802 | 0.964 | 2.942 | 0.998 | 2.254 | 0.988 |
| 4 | 5.477 | 1.000 | 6.091 | 1.000 | 7.211 | 1.000 |

| b) Pesaran (2007) CIPS test: Constant and deterministic trend | | | | | | |
|---|------------|--------------|----------|--------------|-----------|--------------|
| Lags | <i>lis</i> | (<i>p</i>) | <i>q</i> | (<i>p</i>) | <i>rp</i> | (<i>p</i>) |
| 0 | 1.044 | 0.852 | -2.068 | 0.019 | 2.483 | 0.993 |
| 1 | 0.390 | 0.652 | -1.628 | 0.052 | 2.052 | 0.980 |
| 2 | -0.033 | 0.487 | 1.304 | 0.904 | 0.998 | 0.841 |
| 3 | 5.280 | 1.000 | 6.785 | 1.000 | 6.006 | 1.000 |
| 4 | 8.090 | 1.000 | 8.949 | 1.000 | 9.127 | 1.000 |

Notes: Pesaran (2007) CIPS test values are obtained from the standardised Z-tbar statistic. H_0 = nonstationarity. Lags indicates the number of lags included in the ADF regression.

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Intuition

$$y_{it} = \beta_i x_{it} + \phi_i f_t + \psi_i + \varepsilon_{it}, \quad (\text{A.1})$$

$$\bar{y}_t = \bar{\beta} \bar{x}_t + \bar{\phi} f_t + \bar{\psi}, \quad \bar{\varepsilon}_t \rightarrow 0 \quad \text{as} \quad N \rightarrow \infty \quad (\text{A.2})$$

$$f_t = \bar{\phi}^{-1} (\bar{y}_t - \bar{\psi} - \bar{\beta} \bar{x}_t) \quad (\text{A.3})$$

Substitution for f_t in equation (A.1):

$$y_{it} = \beta_i x_{it} + \phi_i \bar{\phi}^{-1} (\bar{y}_t - \bar{\psi} - \bar{\beta} \bar{x}_t) + \varepsilon_{it}, \quad (\text{A.4})$$

$$y_{it} = \beta_i x_{it} + \Pi_{1i} \bar{y}_t + \Pi_{2i} \bar{x}_t + \Pi_{3i} + \varepsilon_{it} \quad (\text{A.5})$$

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Table A.4: Error Correction Model: WID Q

| | [1] 2FE | [2] CCEP | [3] MG | [4] CMG | [5] CMGt | [6] CMGt1 | [7] CMGt2 |
|--------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| lis_{t-1} | -0.136 (0.047)*** | -0.192 (0.068)*** | -0.461 (0.111)*** | -0.442 (0.112)*** | -0.579 (0.155)*** | -0.714 (0.181)*** | -0.958 (0.337)*** |
| qt_{t-1} | -0.001 (0.009) | -0.003 (0.012) | -0.039 (0.036) | -0.001 (0.007) | -0.05 (0.029)* | -0.066 (0.038)* | -0.135 (0.078)* |
| rpt_{t-1} | 0.043 (0.032) | 0.075 (0.055) | 0.151 (0.108) | 0.108 (0.093) | -0.019 (0.054) | -0.044 (0.124) | 0.296 (0.461) |
| Δq | -0.039 (0.018)** | -0.052 (0.022)** | -0.061 (0.039) | -0.043 (0.018)** | -0.042 (0.010)*** | -0.049 (0.020)** | -0.091 (0.036)** |
| Δrp | 0.088 (0.076) | 0.078 (0.080) | -0.062 (0.104) | 0.038 (0.077) | 0.02 (0.075) | 0.158 (0.054)*** | 0.094 (0.297) |
| t | | | 0.001 (0.001) | | 0.002 (0.003) | 0.001 (0.004) | -0.002 (0.004) |
| Constant | -0.066 (0.034)* | | -0.349 (0.082)*** | 0.048 (0.129) | 0.143 (0.130) | 0.273 (0.179) | 0.181 (0.253) |
| Number of id | 9 | 9 | 9 | 7 | 7 | 7 | 6 |
| Observations | 199 | 199 | 199 | 175 | 175 | 171 | 149 |
| R-squared | 0.51 | 0.75 | | | | | |
| RMSE | 0.0124 | 0.0098 | 0.0106 | 0.0067 | 0.0061 | 0.0051 | 0.0039 |
| Trends | | | 0.22 | | 0.43 | 0.14 | 0 |
| $lr-q$ | -0.0052 | -0.0164 | -0.0847 | -0.0011 | -0.0863 | -0.0919 | -0.1404 |
| $se-q$ | 0.065 | 0.0599 | 0.0799 | 0.0149 | 0.0556 | 0.0576 | 0.0949 |
| $lr-rp$ | 0.3149 | 0.3911 | 0.3266 | 0.2434 | -0.0324 | -0.062 | 0.3092 |
| $se-rp$ | 0.2716 | 0.3177 | 0.247 | 0.2199 | 0.0938 | 0.1747 | 0.4931 |
| CD test | -3.8732 | -2.7485 | 3.7987 | -2.0474 | -2.347 | -2.4567 | -1.9305 |
| Abs Corr | 0.2378 | 0.2169 | 0.3325 | 0.2104 | 0.2141 | 0.2757 | 0.2229 |
| Int | I(0) | I(0) | I(0) | I(0) | I(0) | I(0) | I(0) |

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Weak Exogeneity Test

We estimate an informal causality test based on the Granger Representation Theorem (GRT), which states that cointegrated series can be represented in the form of an ECM, which in our case is:

$$\begin{aligned}\Delta lis_{it} &= \alpha_{1i} + \lambda_{11}\hat{u}_{i,t-j} + \sum_{j=1}^K \phi_{11ij}lis_{i,t-j} + \sum_{j=1}^K \phi_{12ij}q_{i,t-j} + \sum_{j=1}^K \phi_{13ij}rpi_{i,t-j} + \epsilon_{1it}, \\ \Delta q_{it} &= \alpha_{2i} + \lambda_{21}\hat{u}_{i,t-j} + \sum_{j=1}^K \phi_{21ij}lis_{i,t-j} + \sum_{j=1}^K \phi_{22ij}q_{i,t-j} + \sum_{j=1}^K \phi_{23ij}rpi_{i,t-j} + \epsilon_{2it}, \\ \Delta rpi_{it} &= \alpha_{3i} + \lambda_{31}\hat{u}_{i,t-j} + \sum_{j=1}^K \phi_{31ij}lis_{i,t-j} + \sum_{j=1}^K \phi_{32ij}q_{i,t-j} + \sum_{j=1}^K \phi_{33ij}rpi_{i,t-j} + \epsilon_{3it},\end{aligned}$$

where $\hat{u}_{it} = lis_{it} - \hat{\beta}_{1i}q_{it} + \hat{\beta}_{2i}rpi_{it}$ is the disequilibrium term.

Table A.5: Weak exogeneity test

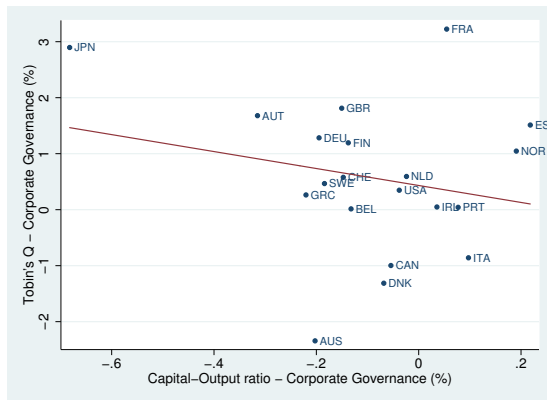
| Model | | no CA | | | CA | | |
|-------|----------------|------------|----------|-----------|------------|----------|-----------|
| | | <i>lis</i> | <i>q</i> | <i>rp</i> | <i>lis</i> | <i>q</i> | <i>rp</i> |
| MG | Avg. λ | -0.52 | -0.45 | 0.02 | -0.50 | -0.41 | -0.04 |
| | ρ | 0.00 | 0.03* | 0.48 | 0.00 | 0.21 | 0.60 |
| CMG | Avg. λ | -0.57 | -0.40 | -0.01 | -0.51 | -0.54 | 0.00 |
| | ρ | 0.00 | 0.15 | 0.83 | 0.00 | 0.18 | 0.94 |
| CMGt | Avg. λ | -0.75 | -0.65 | 0.00 | -0.69 | -0.74 | -0.04 |
| | ρ | 0.00 | 0.01* | 0.98 | 0.00 | 0.12 | 0.72 |
| CMG1 | Avg. λ | -0.59 | -0.23 | 0.04 | -0.51 | -0.58 | 0.03 |
| | ρ | 0.00 | 0.52 | 0.24 | 0.00 | 0.13 | 0.61 |
| CMGt1 | Avg. λ | -0.77 | -0.12 | 0.06 | -0.75 | -0.60 | 0.05 |
| | ρ | 0.00 | 0.75 | 0.19 | 0.00 | 0.19 | 0.38 |
| CMG2 | Avg. λ | -0.73 | -0.42 | -0.07 | -0.64 | -1.04 | -0.05 |
| | ρ | 0.00 | 0.32 | 0.09* | 0.00 | 0.04* | 0.56 |
| CMGt2 | Avg. λ | -0.93 | -0.46 | 0.06 | -0.82 | -1.20 | 0.05 |
| | ρ | 0.00 | 0.29 | 0.25 | 0.00 | 0.01* | 0.44 |

Notes: Avg. λ shows the robust mean coefficient for the disequilibrium term on the ECM. Asterisks highlight cases which do not support a causality relationship for our analysis.

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Corporate Governance (IV): 2002-2007

Figure A.2: Tobin's Q, Capital-Output ratio and Corporate Governance: 2002-2007



Notes: Own calculations obtained from $\ln(X_t) = \alpha_0 + \alpha_1 GOV_t + \epsilon_t$, where X represents the Tobin's Q and the capital-output ratio in the vertical and the horizontal axis respectively. GOV is the corporate governance index, and ϵ is a classic disturbance term. Both axis show the coefficient α_1 in %. Both equations are constraint to have the same number of observations. The scatter plot is obtained after excluding outliers. An outlier is defined as an observation with a weight of 0 after using the *rreg* command in STATA.